

Reply to: “Neutral tumor evolution?”

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We thank Tarabichi and colleagues for the constructive criticism of our Williams *et al.* 2016¹ work. Their critique has four main points that we address below.

1. Impact of clonal copy number alterations

In Williams *et al.* 2016¹, we assessed the cumulative variant allele frequency (VAF) distribution $M(f)$ over the frequency range $f=[0.12,0.24]$ to restrict our analysis to subclonal variants within a range applicable to the diverse datasets we considered. Tarabichi and colleagues note that tumours with a tetraploid genome will have an additional ‘peak’ of clonal mutations at $f\sim 0.25$ (mutations in a single allele, Supplementary Figure 1A), thus causing incorrect rejection of neutrality (Supplementary Figure 1B). The integration range that we chose was based on a triploid tumour with read depth of 100X, giving an upper threshold of 0.26 (see Supplementary Methods). Although this is suitable for most cases, it is not for a tetraploid tumour, suggesting that the number of tumours consistent with neutral evolution could be larger than we reported. In Supplementary Figure 1C we show how this problem can be avoided by adjusting the range for tetraploid tumours.

We do acknowledge that the $1/f$ integration method is more accurate when applied to the whole VAF spectrum of subclonal mutations only. Moreover, we have recently developed a Bayesian model selection framework that compares the neutral model against models with selection, using the entire VAF distribution². We care to stress however, that the majority of cancers analysed in our original manuscript were *not neutral*, and showed signs of subclonal selection.

2. Interpretation of the $1/f$ statistical test

Tarabachi and colleagues correctly note that failing to reject the null is not necessarily evidence for it. While this is true, hypothesis-driven evolutionary analysis of cancer genomic data requires a sensible null. Analysing data without knowing what to expect in the simplest scenarios may lead to wrong conclusions, as we have highlighted³. We have proposed neutrality, the null model of molecular evolution⁴, as a sufficient explanation of the available data from a proportion of tumours. The test we applied quantifies the deviation from the null distribution in terms of a change in the model parameter ($s=0$ vs $s>0$). This structure arises from a frequentist approach and arguments for setting up the test in any other way are arbitrary and impractical. There are an infinite number of models of selection, some producing vanishingly small and so unmeasurable deviations from neutrality (e.g. weak selection), while others are biologically unrealistic (e.g. every mutation is a driver, constant population size in cancer). This is why in molecular evolution neutrality is a suitable null⁴. In our view, selection is arguably the most important force in cancer, but a sensible null model avoids over-interpreting data.

Tarabichi and colleagues state that the $M(f) \sim 1/f$ deterministic solution we reported in our manuscript (Eq.7) relies on the strong assumption of synchronous cell divisions. That is not the case: Eq.7 is the convergent solution of a continuous-time stochastic branching process for large number of cells⁵. Simulations based on Gillespie algorithm that explicitly model asynchronous cell divisions agree with the solutions of the stochastic branching process. Tarabichi *et al.* also state that simulating stochastic processes is more realistic. This is indeed what we did in our original manuscript (Fig S9-S11 in Williams *et al.* 2016), demonstrating the convergence to the deterministic solution. A comprehensive analysis of the underlying stochastic Luria-Delbrück model shows that the scaling behaviour of the stochastic branching process ($1/f$ tail) remains unchanged even in the explicit presence of stochastic cell death⁶. In Tarabichi's letter Figure 1b, the claim that a stochastic neutral model does not imply $1/f$ is therefore incorrect, as also demonstrated by others before us⁵⁻⁷. Moreover, in the beginning of their letter Tarabichi and colleagues argue that the $1/f$ tail is solely due to drift. This is true only in populations of constant size. In exponentially growing populations that start from a single cell such as cancer, the $1/f$ subclonal tail emerges instead from somatic mutations acquired at different times during growth.

3. Insights from simulated tumours

Tarabichi and colleagues use a stochastic branching process, similar to our previous implementation^{1,2}, to generate synthetic genomic data and test our method. In their Figure 1, Tarabichi *et al.* present a synthetic analysis of the $1/f$ test using the analytical deterministic solution (Figure 1A) and stochastic simulations (Figure 1B). In both analyses, a new subclone is introduced

at a certain fixed time point. First, we note that Tarabichi's Figure 1a is different from 1b, which contradicts mathematical theory^{1,5-8}.

The parameters used in Tarabichi's simulations are also rather extreme. In their model the emergence of a subclonal driver increases the net-growth rate ($adv_{subclone}$, selective advantage), and modifies the mutation rate at the same time. Curiously, the mutation rate can decrease by a factor of 8, or increase by a factor 100. A rate of 1024 new mutations per cell division (Tarabichi's Figure 1a, x-axis, $\mu_{subclone}=2^{10}$) pertains only to a very small set of colorectal and uterine cancers with POLE/POLD mutations. A POLE subclone arising within a POLE wild-type background is a very rare event⁹. Thus, we urge caution when considering the implications of the parameters at the extremities of the range considered by Tarabichi *et al.*

To address the authors' criticism, we have reproduced their Figure 1b with our stochastic branching process (Figure 1A, see Supplementary Methods). We systematically found that when a subclone with $f_{subclone} \geq 10\%$ is selected (selective advantage $adv_{subclone} > 0.5$), the 1/f test identifies it correctly (neutrality is rejected; top left quadrant of Figure 1A, example in 1B). When the new subclone is instead very small (weak selection), the 1/f test fails to reject neutrality ($adv_{subclone} < 0.5$, bottom half of Figure 1A, example in 1C). Figure 1D illustrates the relationship between the selective advantage and the subclone cell fraction in the final tumour, highlighting the issue of the limit of detectability (LOD). We have subsequently quantified this effect², identifying a 'wedge of selection' that describes the detectability problem in cancer genomic data at current resolution. If subclonal selection does not significantly change the clonal composition of the tumour, the signature of neutral growth ('1/f tail') will still dominate the detectable VAF spectrum (bottom part of Figure 1A).

Notably, for a hypermutant subclone with strong selective advantage ($\mu_{subclone} \geq 64$, top right-hand side of Figure 1A), the analysis showed a massive 1/f tail containing thousands of the subclone's private mutations. These mutations dominate the entire VAF distribution and obscure the underlying subclonal structure (Figure 1A, example in 1E). It is not surprising that our test, or any other test, would struggle to detect any subclonal cluster or deviation from 1/f in these cases. Curiously, for moderate values of selection ($adv_{subclone} \sim 0.5$), a change in mutation rate from normal to hypermutant could be detected, leading to rejection of neutrality (mid-right area in Figure 1A, example in 1F, see also Fig S11H in Williams *et al.* 2016). For weak selection and a hypermutator subclone, the new subclone did not reach a detectable size and therefore neutrality could not be rejected (Figure 1C).

Importantly, we note that the lack of discriminatory power in these peculiar scenarios does not depend on our method but it is largely due to minimal signal in the data. To demonstrate this, we

compare our method with the extended integration range (Figure 1G) to DPclust, a method to detect subclones based on Dirichlet Process clustering¹⁰ proposed by some of the authors of the letter (see Supplementary Methods). Even under optimal circumstances (strong selection), the sensitivity of DPclust is suboptimal in the vast majority of cases (Figure 1H).

We are pleased however that Tarabichi and colleagues confirmed that 1/f tails are pervasive in cancer genomic data. Neutral tails are a simple consequence of clonal growth, and appear within each individual clone during its expansion^{1,2}.

4. Analysis of subclonal selection using dN/dS ratios

Using a test inspired by the classical dN/dS method, Tarabichi *et al.* argue to find evidence of subclonal selection in tumors classified by our test as neutral. Specifically, the authors pool together subclonal mutations in known cancer genes from multiple patients, and calculate a dN/dS value for neutral and non-neutral groups. Their criticism is that subclonal mutations in the neutral group should lack evidence of selection (dN/dS~1).

First, we note that it is wrong to draw conclusions about individual samples from such a population-level statistic. Instead, discrepancies between the dN/dS value and our 1/f test results could simply mean that our method may have misclassified one, or more, patients. To investigate this, we reproduced their analysis using the same dN/dS method¹¹ and measured global dN/dS for 369 driver genes¹¹ in colorectal and gastric cancers analysed in our original manuscript¹. Tarabichi's TCGA pan-cancer analysis used CAVEMAN calls that are not publicly available, so instead we reanalysed the pan-cancer TCGA variant calls publicly available from the GDC data portal (see Supplementary Methods). We found that in all three cohorts, dN/dS of subclonal missense mutations in neutral-classified tumours was not significantly different from 1, thus confirming our findings (Figure 2A-C, missense mutations on the left, blue bars).

Interestingly, we found a small group of neutrally classified patients with an unusually high number of subclonal nonsense mutations in putative driver genes. 1/57 gastric cases (1.7%) and 11/290 (3.8%) of pan-cancer cases had 3 or more subclonal nonsense mutations. We manually examined these patients (Supplementary Figure 2) and found that these were affected by clonal mutations 'bleeding' into the subclonal integration range, misclassification caused by ploidy measurement errors, and possibly the presence of selected subclones hidden underneath 1/f tails. We care to stress that these tumours had been classified with our original limited integration range. After removing patients with 3 or more nonsense mutations (e.g. leaving 96.2% of putatively neutral cases in the pan-cancer cohort), the dN/dS value for nonsense was not significantly different from the neutral background (Figure 2C, dN/dS=1.44, p=0.32). We demonstrate this by generating

dN/dS values for 'control sets' (background) of passenger genes using bootstrapping of 1,000 random sets of 198 non-driver genes¹² as well as neutral genes (Figure 2D). This analysis indicates a systematic positive bias for the estimation of dN/dS. This could be due to public GDC calls being depleted of synonymous somatic mutations present in dbSNP, skewing dN/dS values as mentioned in Martincorena *et al.* 2017. Interestingly, subclonal dN/dS values were consistently higher in non-neutral versus neutral cases, although this was not significant (Figure 2D).

We care to highlight that dN/dS analysis at the cohort-level combines mutations from different patients, whereas the neutrality test is patient-specific. Even a single misclassified patient carrying multiple nonsense mutations in driver genes significantly alters the dN/dS value of a whole cohort. While dN/dS can reveal an excess or depletion of mutations in a cohort, Tarabichi *et al.*'s analysis cannot differentiate whether this is coming from one or multiple patients.

Summary

We thank Tarabichi and colleagues for providing some valid constructive criticism of our original manuscript. In our assessment of their critique however, our original conclusion remains valid: that neutral evolution provides an entirely adequate description of the pattern of intra-tumour heterogeneity that has been observed to date across many tumours. We are also grateful to Tarabichi and co-authors because they led us to the finding that VAF distribution analyses applied to single patients, like our neutrality test, can be carefully combined with cohort-level statistics like dN/dS to increase the power to discriminate between neutral dynamics and selection in cancer.

Contributions

TH performed simulation and bioinformatic analyses. LZ performed dN/dS analysis. MW performed additional simulation analysis. BW performed mathematical analysis. GC provided expertise on clustering analysis. CB, TAG and AS conceived and led the study. All authors addressed the criticism and wrote the response.

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Competing financial interests

We declare no competing financial interests.

Figures

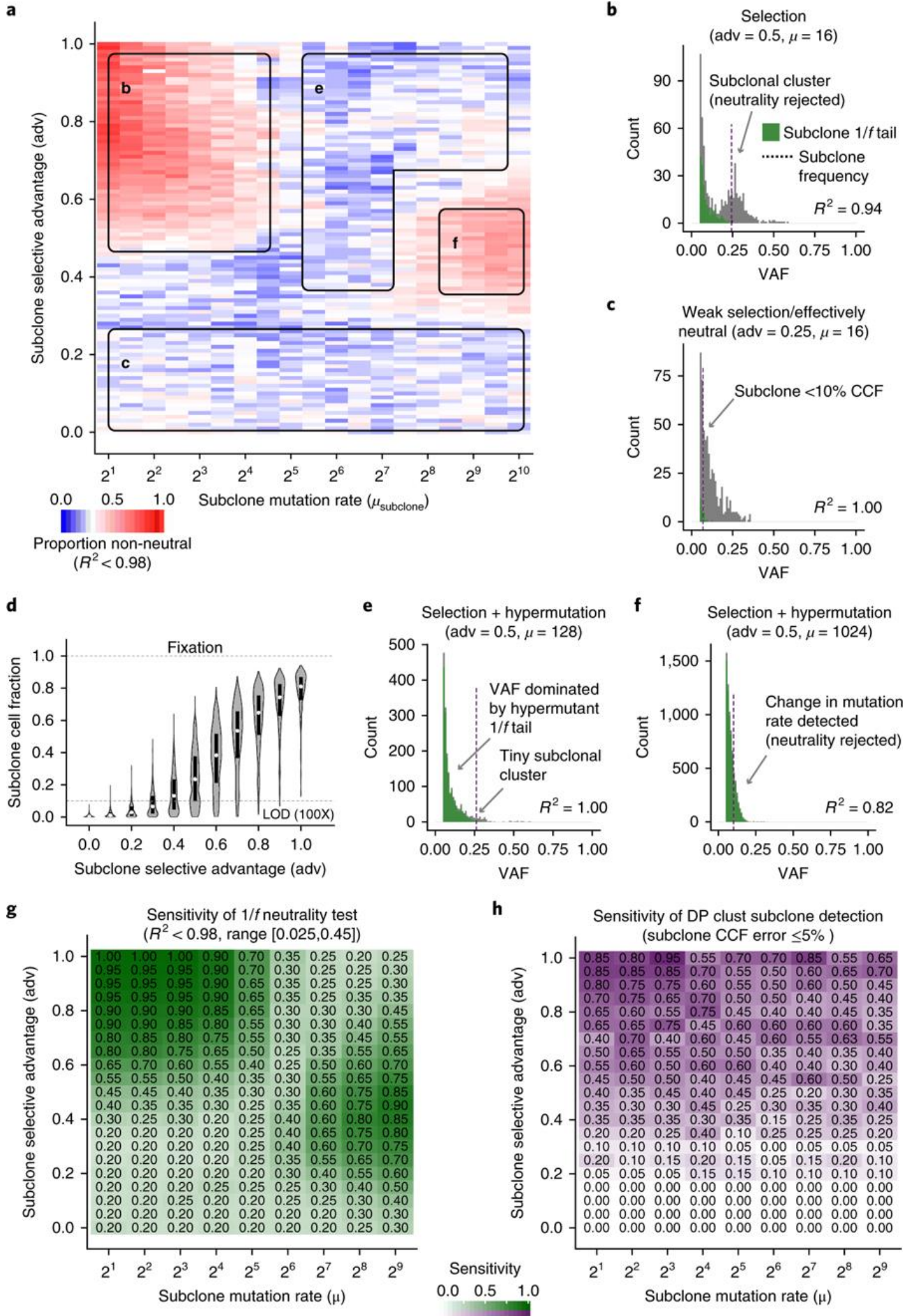


Figure 1. Insights from stochastic simulations of cancer growth. (A) Heatmap recapitulating Tarabichi's Figure 1b with same parameter set and showing proportion of simulations where neutrality was rejected (200 cases per parameter combination). **(B)** Example VAF distribution with a detectable subclonal cluster (dashed line indicates subclone frequency). The $1/f$ test rejects neutrality in favour of selection (R^2 reported). **(C)** Example VAF distribution with a weakly-selected subclone that remains below the limit of detection (100X depth). **(D)** Subclone cell fraction in the final tumour as a function of fitness advantage, for $\text{adv}_{\text{subclone}} < 0.5$ the subclone rarely reaches detectable size of $\sim 10\%$ cell fraction (assuming 100x depth). **(E)** Example VAF distribution for a subclone with selective advantage and, at the same time, high mutation rate. **(F)** Example VAF distribution for a selected and extreme mutator subclone. **(G)** Sensitivity of the $1/f$ test applied to subclonal mutations in the extended range of $\text{VAF}=[0.025,0.45]$ from the simulations in panel A, numbers report proportion of cases where neutrality was rejected ($R^2 < 0.98$). **(H)** Sensitivity of subclone detection of DPclust, a Dirichlet subclonal clustering method, when applied to the same simulated data. Numbers report proportion of cases (20 cases per combination) where the correct subclone has been identified within a 5% CCF error with respect to true position.

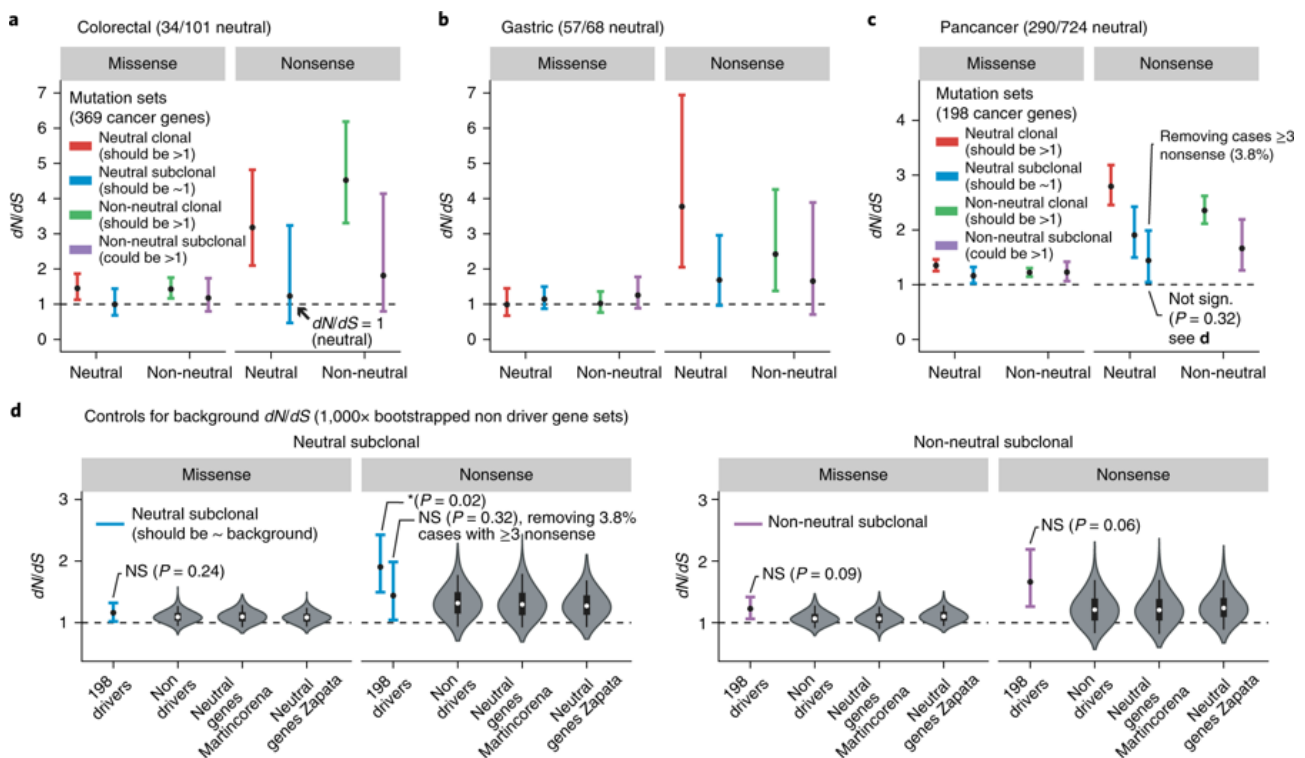


Figure 2. Detecting subclonal selection with dN/dS analysis. dN/dS analysis using Martincorena *et al.* 2017 method applied to colorectal cancers **(A)**, gastric cancers from ref¹³ analysed in Williams *et al.* 2016 **(B)**, and TCGA pan-cancer cases using newly available GDC calls to reproduce Tarabichi's dN/dS analysis **(C)**. Cancers were classified as neutral or non-neutral using the $1/f$ test, and the dN/dS values of were calculated over pooled variants from each group

(split between clonal/subclonal and missense/nonsense). **(D)** Comparison of the dN/dS estimates obtained for the 198 driver genes (black dot: point estimate, error bars: 95% CI) with the distribution of 1,000 random subsets from three control sets of non-driver genes demonstrates a general positive bias of estimated dN/dS values (white dot: median, box: interquartile range, whiskers: 90% prediction interval). After removing 3.8% of pan-cancer cases with ≥ 3 subclonal nonsense mutations in driver genes, both missense and nonsense dN/dS in neutral cancers were not significantly different from the neutral expectation.

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