Smoking and risk of breast cancer in the Generations Study cohort

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ABSTRACT

Background: Plausible biological reasons exist why smoking could affect breast cancer risk but epidemiological evidence is inconsistent.

Methods: We used serial questionnaire information from the Generations Study cohort (United Kingdom) to estimate hazard ratios (HRs) for breast cancer in relation to smoking adjusted for potentially confounding factors including alcohol intake.

Results: Among 102,927 women recruited 2003–2013, with 7.7 years average follow-up, 1815 developed invasive breast cancer. The HR (reference group: never smoker) was 1.14 (95% confidence interval (CI): 1.03–1.25; \(P=0.010\)) for ever-smoking, 1.24 (95% CI: 1.08–1.43; \(P=0.002\)) for starting smoking at ages <17 years, and 1.23 (1.07–1.41; \(P=0.004\)) for starting smoking 1–4 years after menarche. Breast cancer risk was not statistically associated with interval from initiation of smoking to first birth (\(P\)-trend=0.97). Women with a family history of breast cancer (ever smoker vs never smoker HR=1.35; 95% CI: 1.12–1.62; \(P=0.002\)) had significantly larger HR in relation to ever-smoking (interaction: \(P=0.039\)) than women without (ever smoker vs never smoker HR=1.07; 95% CI: 0.96–1.20; \(P=0.22\)); the interaction was prominent for age started smoking (\(P=0.003\)) and starting smoking relative to age at menarche (\(P=0.0001\)).

Conclusions: Smoking was associated with a modest but significantly increased risk of breast cancer, particularly among women who started smoking at adolescent or peri-menarcheal ages. The relative risk of breast cancer associated with smoking was greater for women with a family history of the disease.

KEYWORDS:

smoking, breast neoplasms, cohort studies

BACKGROUND
The carcinogenic potential of tobacco smoke is unarguable [1, 2] and there are plausible biological reasons why smoking could affect breast cancer risk [2-5]. Reviews of the association between cigarette smoking and breast cancer up to 2004 did not, however, generally find conclusive evidence for a causal relationship in humans [5-7]. More recent epidemiological analyses have reported modest raised risks with current [8-19] or former [8-15, 20] smoking, but questions remain about the extent to which this association is a consequence of confounding by alcohol use, whether risk is increased if smoking starts in adolescence or before first childbirth, and whether risk is modified by family history of breast cancer [1, 2]. We therefore examined risk of invasive breast cancer in relation to smoking in a large cohort study using detailed questionnaire information at recruitment and during follow-up, with adjustment for alcohol consumption and other potentially confounding factors.

**METHODS**

The Generations Study is a cohort study of over 113,700 women aged 16 or older from the United Kingdom, from whom questionnaire information and informed consent was gained at recruitment since 2003 [21]. Initial recruits to the cohort were from women involved in the breast cancer charity that funded the study, and women who responded to publicity about the study. Women who joined the study were asked to nominate female friends and family members, who were then contacted about joining the study. This referral method continued with subsequent recruits [21]. The first follow-up questionnaire (2½ years after recruitment) was completed by 99% of non-deceased participants, a second (six years after recruitment) by 96%, and a third (9½ years after recruitment) by 94% (of those recruited long enough ago to have entered this round of follow-up). The study was approved by the South East Multi-Centre Research Ethics Committee.

Breast and other cancers occurring in the cohort were identified from recruitment and follow-up questionnaires, spontaneous reports to the study centre, and from ‘flagging’ (see below) for those lost to questionnaire follow-up. Confirmation of diagnosis was obtained from cancer
registries in the United Kingdom, ‘flagging’ at the National Health Service Central Registers (virtually complete registers of the populations of England and Wales, and of Scotland, to which study participants can be linked and on which deaths, cancer registrations, and emigrations are ‘flagged’ and then periodically reported to authorized medical researchers), pathology reports, and correspondence with patients’ general practitioners.

Information on risk factors for breast cancer was obtained from recruitment and follow-up questionnaires. In relation to smoking, women were asked if they had “ever smoked regularly (i.e. most days for at least 6 months)”, if they still smoked regularly, age started and stopped, and number of cigarettes smoked per day at different periods of their lives (during ages: 16–24, 25–49, 50+ years). For analysis, we defined the period of ‘current smoking’ to include both current smokers and the year immediately after stopping, to avoid potential ‘reverse-causation’ bias from women who may have stopped smoking during the work-up to a formal breast cancer diagnosis. For alcohol use we asked women if they had been a regular drinker “in the sense of drinking at least one glass of alcohol per week on average”, ages started and stopped, and quantity consumed at different periods of life (during ages: 18–24, 25–49, 50+ years). We converted the quantity of alcohol consumed at each period of life into daily grams of alcohol. We split into three groups the women who reported current drinking (<60g/day, 60+g/day, and amount unknown), and we classified women who had reported stopping drinking as former drinkers. For some women we did not know their current drinking status during follow-up, but we knew they had consumed alcohol in the past and these women were classified as ‘ever-drinkers’. Because we had collected ages or dates at which certain events or changes in lifestyle occurred we were able to update smoking status, alcohol use, parity, oral contraceptive (OC) use, menopausal hormone therapy (MHT) use, and menopausal status, at the ages these episodes occurred through to the second follow-up questionnaire. We updated duration of smoking for current smokers, and time since cessation for former smokers, in yearly increments, using smoking start and stop ages from the recruitment and second follow-up.
questionnaire. We updated cigarettes smoked per day, pack-years smoked, alcohol consumption, and post-menopausal body mass index (BMI), at the date of the second follow-up questionnaire.

Statistical analysis

The current analytic cohort is based on all women who were recruited to the study during June 2003–December 2013 without prior invasive or in-situ breast cancer or other malignant cancer (except non-melanoma skin cancer), or prior mastectomy. The recruitment cut-off at December 2013 was selected because at the time of analysis the second follow-up was practically complete for this group of recruits, two-thirds of the cohort had reached the third follow-up, and we had ‘flagging’ information to June 2017. Women entered risk at their date of recruitment and were censored at the earliest date of: invasive breast cancer or in-situ breast cancer; other malignancy (except non-melanoma skin cancer); death; most recent follow-up questionnaire (depending on date of recruitment) if completed, or the date most recent follow-up questionnaire was due if cancer and vital status was known from ‘flagging’; or previously completed questionnaire if lost to follow-up. We censored follow-up at in-situ breast cancer or other malignancy because we reasoned that if smoking is related to risk of in-situ breast cancer or other malignancy, and ensuing treatments or their consequences alter risk of subsequent invasive breast cancer, including subsequent follow-up may obscure associations between smoking and invasive breast cancer.

Left-truncated and right censored Cox proportional hazards regression [22] using attained age as the implicit time scale was used to estimate hazard ratios (HR) and 95% confidence intervals (CI) for smoking and risk of first invasive breast cancer. We adjusted for: time since recruitment to cohort (0, 1–2, 3+ years); birth cohort (1908–39, 1940–49, 1950–59, 1960–69, 1970–96); benign breast disease (yes, no); family history of breast cancer in 1st degree relatives (yes, no); socio-economic score (ACORN score as trend, or missing indicator); age at menarche (trend, or missing indicator); age at first pregnancy (trend, or missing indicator); parity (trend, or missing indicator); duration of breastfeeding (trend, or missing indicator); current OC use during follow-up, before
menopause (yes, no); alcohol consumption (trend for current drinker 1–<60g/day, indicator variables for never regular, current drinker 60+g/day, past drinker, drinker with unknown details); physical activity (log(metabolic equivalent) trend, missing indicator); pre-menopausal BMI at age 20 years (trend, or missing indicator); post-menopausal BMI (trend, or missing indicator); MHT use (never used, ex-user, current estrogen only user, current estrogen plus progestogen user, current user of other types, missing indicator); menopausal status (pre- or post-menopausal) and age at menopause (trend, or missing indicator). BMI was used to create two separate variables: pre-menopausal BMI (potentially available for all women) and post-menopausal BMI (only available at post-menopausal ages). We used BMI at age 20 to represent pre-menopausal BMI. Separately, if a woman was post-menopausal at entry to the cohort we used her BMI at entry for her post-menopausal BMI (and if she was pre-menopausal at this time, post-menopausal BMI was unknown). If a woman was post-menopausal at the time of the follow-up questionnaire we updated from this point in time her post-menopausal BMI with the value from this follow-up questionnaire. Statistical trends were evaluated using continuous values, except for duration and time since cessation of smoking which were based on discrete time-varying annually updated values. For trend analyses where there was an unexposed group (e.g. never smokers in analyses of smoking duration) the unexposed group was not assigned a zero magnitude but was treated as a separate categorical term, as was any missing value group. In particular we adjusted our analyses of smoking and breast cancer for alcohol using daily current alcohol consumption as a continuous measure if within the range 1–<60g/day, and categorical terms for non-drinkers, for those with consumption 60+g/day (because we did not want a minority of women who reported very high consumption to influence unduly the trend with daily consumption), past drinkers, and those for whom details of consumption were missing, by fitting appropriate interaction terms in the Cox regression model. Heterogeneity in HRs by sub-type of breast cancer defined by estrogen receptor (ER) status or morphology was assessed using a data augmentation method [23] and Wald chi-square tests [24]. All statistical tests were two-sided and analyses were conducted using Stata/IC version 14.0 [25].
RESULTS

During 2003–2013 a recruitment questionnaire was completed by 102,940 women who had no previous invasive or in-situ breast cancer or other malignancy (except non-melanoma skin cancer). At censoring date 1.1% of women had died. Of the remainder, cancer and vital status was known for 96.5% who had completed the relevant follow-up questionnaire, and a further 2.4% from ‘flagging’ at the National Health Service Central Registers. The remaining 1.1% were lost to follow-up at an earlier date. Thirteen women (including one with breast cancer) were excluded from subsequent analyses because of self-contradictory information for parity or smoking, leaving 102,927.

Table 1 presents descriptive characteristics at recruitment of the cohort eligible for analysis.

The median age at recruitment was 47 years (Inter-Quartile Range (IQR): 36–57). A majority of participants, 64.1%, reported never smoking but only 10.3% were never-regular consumers of alcohol. In relation to alcohol consumption, 12.5% of never-smokers were non-drinkers in contrast to 6.4% of ever-smokers. Among those who reported drinking <60g/day the median alcohol consumption (g/day) was 14.2 (IQR: 8.7–22.1) among never smokers and 19.0 (IQR: 11.9–29.2) among ever smokers. Supplementary Table 1 provides further descriptive characteristics of the cohort in relation to age started smoking, thelarche, parity, menopausal status, and BMI.

During 788,361 person-years (median 6.6 years; mean 7.7 years) of follow-up 1815 invasive breast cancers were diagnosed, of which 1813 were confirmed through national cancer registration or medical records, and the remaining two were self-reported with treatments that imply breast cancer. ER-status data were available for 99.3%, and of these 83.7% were ER-positive. Invasive ductal carcinoma accounted for 78.8%, and lobular 16.4%, of tumours. Further descriptive characteristics of the breast cancer cases are given in Supplementary Table 2.
The HR for invasive breast cancer in relation to ever smoking was 1.17 (95% CI: 1.07–1.29; \( P=0.0009 \)) when adjusted only for attained age, 1.13 (95% CI: 1.03–1.24; \( P=0.012 \)) when also adjusted for alcohol consumption, and 1.14 (95% CI: 1.03–1.25; \( P=0.010 \)) when further adjusted for other potentially confounding variables (see Methods and Table 2). All subsequent results are adjusted for attained age, alcohol consumption and the potentially confounding variables, unless otherwise stated.

Table 2 presents results for breast cancer overall and by ER status. The HR for ever-smoking was raised for ER-positive (HR=1.12; 95% CI: 1.01–1.24; \( P=0.035 \)) and ER-negative (HR=1.25; 95% CI: 0.99–1.58; \( P=0.063 \)) breast cancer, and the difference between the HRs was not significant (\( P=0.40 \)). Breast cancer risk increased significantly with number of cigarettes smoked per day for all breast cancer (\( P\text{-trend}=0.0060 \)) and for ER-positive tumours (\( P\text{-trend}=0.023 \)). Breast cancer risks were raised significantly after 10+ years duration of smoking (10+ years vs never-smoking: \( P=0.0004 \)). Breast cancer risks did not further rise beyond 10 years duration and because of this non-linear relationship there was no significant linear trend with duration of smoking (\( P\text{-trend}=0.24 \)), nor was there significant heterogeneity in the trend by ER status. Pack-years of smoking was associated with breast cancer risk overall (\( P\text{-trend}=0.0069 \)) and ER-positive breast cancer (\( P\text{-trend}=0.024 \)) but not for ER-negative (\( P\text{-trend}=0.16 \)) tumours; there was no significant heterogeneity of the pack-years trend by ER status (\( P=0.66 \)).

The HR within the year after smoking cessation was 2.68 (95% CI: 1.60–4.46), based on 15 cases, but for reasons described in Methods this risk period was assigned for further analyses to the ‘current-smoker’ group. On this basis risk of breast cancer was raised in current (HR=1.12; 95% CI: 0.89–1.39; \( P=0.34 \)) and former (HR=1.14; 95% CI: 1.03–1.26; \( P=0.011 \)) smokers although only the latter reached statistical significance; there was no significant heterogeneity by ER status. Breast cancer risks were significantly raised within the first 20 years after cessation of smoking and
decreased with greater time since cessation although the trend was not significant ($P_{-trend}=0.071$) and there was no significant heterogeneity in this trend by ER status.

There was significant variation in risk of breast cancer by age at start of smoking (Table 3) ($P_{-}heterogeneity=0.018$; not presented in Table 3). Breast cancer risk was significantly increased if smoking started at ages <17 (HR= 1.24; 95% CI: 1.08–1.43; $P=0.0023$) or 17–19 (HR= 1.15; 95% CI: 1.01–1.31; $P=0.030$) years relative to non-smokers, but not if it started at older ages. The risk was significantly increased for ER-positive, only for smokers starting at ages <17 years, and no significant risk increase was noted for ER-negative breast cancer. When adjusted for pack-years the breast cancer risk for starting smoking at ages <17 years was (HR= 1.12; 95% CI: 0.96–1.32; $P=0.14$), and when adjusted for duration of smoking it was (HR=1.16; 95% CI: 0.96–1.40; $P=0.11$) (not presented in Table 3).

In our questionnaire we asked women only about the amount they smoked per day beginning at age 16; therefore we could not examine smoking intensity at younger ages. There was no significant trend in breast cancer risk, however, in relation to cigarettes smoked per day at ages 16–24 years. Relative to age at menarche, breast cancer risks were highest if smoking started at or before menarche (HR=1.40; 95% CI: 0.98–1.99; $P=0.061$) or 1–4 years after (HR=1.23; 95% CI: 1.07–1.41; $P=0.0040$), with a significant downward trend in breast cancer risk with increasing interval from age at menarche to age at starting smoking ($P=0.031$). A similar pattern was seen for ER-positive, but was less clear for ER-negative, breast cancer. A weaker relationship was seen with age at thelarche (e.g. 1–4 years after thelarche (HR=1.17; 95% CI: 1.00–1.37; $P=0.056$)). When adjusted for pack-years of smoking the HRs for age started smoking 1-4 years after menarche (HR=1.12; 95% CI: 0.96–1.31; $P=0.15$) or thelarche (HR=1.05; 95% CI: 0.88–1.25; $P=0.59$) were attenuated (not presented in Table 3). There was a comparable attenuation after adjusting for duration of smoking. Among parous women there was a significant trend in breast cancer risk with
interval from starting smoking to birth of first child ($P$-trend=0.013); for an interval of 15+ years the HR was 1.46 (95% CI:1.18–1.81; $P$=0.0005). However, these results were not adjusted for age at first child birth and parity (not in Tables), and when we adjusted (as shown in Table 3) there were no significantly raised HRs, or trends for all breast cancer or by ER status.

When analysed by morphological type (Supplementary Table 3) we found significant associations for ductal breast cancer similar to the results for breast cancer overall, and generally non-significant results for lobular breast cancer. There were no significant interactions by morphological type in the risk of breast cancer with smoking.

There was no raised risk of breast cancer with ever-smoking in non-drinkers (HR=0.97; 95% CI: 0.61–1.52; $P$=0.89) but a significantly raised breast cancer risk in those who had ever been drinkers (HR=1.18; 95% CI: 1.07–1.30; $P$=0.0010) although the difference in HRs was not significant ($P$-interaction=0.41) (Table 4). When further stratified by amount of alcohol consumed the HRs for ever-smoking among current drinkers remained raised. Results were similar when we examined breast cancer risk by drinking status for former smokers relative to never smokers (Supplementary Table 4).

We examined further potential risk factor interactions with smoking but found no significant interactions with parity ($P$=0.095) although for nulliparous ever smoking women there was a statistically significantly increased risk of breast cancer ($P$=0.012) (Supplementary Table 5), or menopausal status ($P$=0.73) (Supplementary Table 6), although while the hazard ratio of pre-menopausal ever smokers was somewhat larger than for post-menopausal ever smokers, the former did not reach statistical significance ($P$=0.088), while the latter did ($P$=0.040). Nor did we find significant interactions with birth cohort ($P$=0.092), or BMI at age 20 ($P$=0.55) or post-menopausal ages ($P$=0.26), but we did see a significant interaction with family history of breast cancer ($P$=0.038). There were significant interactions between family history and age at starting smoking ($P$=0.0029).
and starting smoking relative to age at menarche ($P=0.0001$) in relation to risk of breast cancer (Table 5). In particular, among women with a family history of breast cancer, HRs were raised if smoking started at age 20+ years (HR=1.56; 95% CI: 1.17–2.10; $P=0.0028$) or <20 years (HR=1.26; 95% CI: 1.02–1.56; $P=0.029$), and if started 5+ years after menarche (HR=1.53; 95% CI: 1.22–1.91; $P=0.0002$), and we note these were somewhat different to the results among women without a family history of breast cancer.

|TABLE 5 here|

**DISCUSSION**

In the Generations Study cohort we found significant but modestly raised risk of invasive breast cancer in ever and former smokers, in women who smoked more than five cigarettes a day, had 10+ pack-years of use, or had stopped for less than 20 years. Some previous studies have reported similar associations with smoking [8-17, 20], cigarettes per day [9-11, 19], pack-years [9-13, 17-19, 26-29], and cessation [8, 12, 19, 26, 28], but not all studies find these associations [10, 11, 13, 15-17, 19, 20, 29, 30]. We saw significantly raised risk with 10+ years duration of smoking, but no increasing trend beyond 10+ years. Increased risks at long durations (or significant trends) have previously been reported in some studies [8-13, 18-20, 26-28], although some classified non-smokers as smokers with zero duration [12, 20, 26, 28] and this may artefactually produce a significant trend which partly or wholly reflects the difference in risk between non-smokers and smokers (but this may not be the only reason for an association with 20+ years (long duration) of smoking).

We found risk was significantly raised in former smokers, as has been previously reported [8-15, 20]. Risk was also raised with current smoking but the numbers of current smokers in our cohort was small and this result did not reach statistical significance, although some other studies have reported significantly raised risks in this group [8-19]. The raised risks for current and former smokers were similar (HR 1.12 and 1.14) and the confidence intervals overlapped, suggesting, within
our cohort, no material difference between current and former smokers in relation to breast cancer risk.

**Breast cancer sub-types.** We found significant raised risks for ER-positive and ductal breast cancer, which were the most common types in our study, but no significant heterogeneity by ER-status or morphological type of the breast cancer in relation to smoking. The statistical power to examine differences by ER-status or morphology was low in our cohort because of the relative uncommonness of ER-negative and non-ductal type tumours. Some studies have tended to find stronger risks for ER-positive breast cancer [12, 16, 20, 31] but none have found significant interactions and the literature is inconclusive [2]. We observed larger HRs for smoking and pre-menopausal, relative to post-menopausal, breast cancer but the former did not reach statistical significance, and although the literature is variable it does in general suggest a greater relative risk among pre-menopausal women [1, 2]. However, we found no evidence for a significant interaction with menopausal status, similar to other studies [8, 11, 32].

**Confounding by alcohol.** Alcohol consumption was associated with smoking and is itself a known risk factor for breast cancer [7]. We adjusted for alcohol intake and although this reduced the strength of the association between smoking and breast cancer (from HR=1.17 to 1.14) the association remained raised and significant. There is, however, concern that statistical adjustment using self-reported alcohol consumption may not be adequate to control fully for confounding by alcohol [7] so to explore further the extent of potential confounding we stratified by alcohol consumption (Table 4). Within each stratum of consumption (<20g/day, 20–40g/day, and 40–<60g/day) the difference in self-reported alcohol intake between never and ever smokers was ≈1g/day, and we calculate this difference in consumption would be associated with <1% change in relative risk of breast cancer (using the alcohol-breast cancer estimate of relative risk from a large collaborative re-analysis [7]). Within each of these strata it would require ever-smokers to be drinking 20g/day more than never-smokers to produce a difference of ≈15% (similar to the 12–17%
We saw no significant association between smoking and breast cancer risk among non-drinkers, in concordance with a collaborative re-analysis of 43 case-control and 10 cohort studies [7], the American Cancer Society’s CPS II cohort [16], and a subsequent pooled analysis of 14 cohort studies [8]. It is possible there may be synergistic interaction between ever-smoking and alcohol consumption, and risk of breast cancer, although only one study has reported the interaction as statistically significant [8]. There is some precedent to invoke synergism between smoking and alcohol because, for example, there is an established positive interaction between these two exposures and the aetiology of head and neck cancers [33]. However, non-drinking may occur for cultural or religious reasons, or because of underlying illness or other health issues, and in the UK at least non-drinkers are a minority group; therefore this potential interaction could be a reflection of a particular distribution of breast cancer risk factors among non-drinkers (and inadequate control for confounding among drinkers). Conversely, three other cohort studies found significantly raised risk among non-drinkers [18, 26, 29], although in two the increased raised risks were only in subgroups [26, 29].

Smoking in adolescence. Based on epidemiological considerations and animal studies the period from puberty to first birth may represent a window of particular susceptibility to breast cancer [34-37]. At puberty the breast is made up of mainly undifferentiated terminal ductal and lobular structures which animal studies show are sensitive to chemical carcinogenesis [34]. At these young ages ionizing radiation exposure also increases risk of breast cancer [37], especially if exposure is within six months of menarche [38]. We found risk of breast cancer in ever-smokers was greatest if smoking started at ages <17 years, or started at peri-menarcheal or, more weakly, peri-thelarcheal ages. A number of other studies have also found raised risks if smoking started in adolescence [8-13, 16-18, 20, 26, 28, 29, 32] or around menarche [11, 16, 26]. However, when we adjusted for pack-years of smoking the raised risks for starting smoking close to age at menarche or...
thelarche were somewhat attenuated suggesting over-adjustment (because of possible correlation between age starting smoking and pack-years) or confounding by pack-years. Previous studies have not made this adjustment so the relative importance of early initiation or pack-years of use remains unclear.

**Smoking before first childbirth.** Young age at first birth and increasing parity confer long-term protection against breast cancer [34, 35] and animal models point to terminal differentiation of breast tissue at full term pregnancy being important in this process [34-36]. Increased risks have been reported for invasive breast cancer if smoking started before first childbirth [8-11, 16, 17, 20, 26, 28, 29, 32] but we found the association was only significant if we did not adjust for age at first pregnancy. A number of previous studies have adjusted for age at first pregnancy and still found significant associations with interval to first birth [8, 9, 11-13, 16-18, 20, 26, 28, 29] however it is difficult to determine the adequacy of adjustment. For example, in a large pooled analysis of 14 cohort studies there was a strong trend with smoking interval before first birth after adjustment for potential confounders that included age at first birth and number of live births ($P=0.0000002$) whereas after stratification by age at first birth the trends in each strata were weaker ($P=0.12, 0.02,$ and 0.28) [8], which is suggestive of confounding.

**Interaction with family history.** We found the association between smoking and breast cancer was significantly larger among women with a family history of the disease than those without. Five previous studies have reported on this interaction with family history. Two studies reported no significant interaction but did not present stratified results so we cannot determine if the direction of interaction support or contradict our finding [16, 19]. Three studies reported significant interactions, with one showing increased breast cancer risk with smoking only among those with a positive family history [39], whereas two found breast cancer risk was raised only among those with no family history [15, 18]. Increased risk of breast cancer with smoking has also been seen in some [40, 41], but not all (see review [1] and a large meta-analysis [41]), studies of BRCA1/2 carriers (or by proxy, women with three or more first degree relatives with breast or
ovarian cancer[42]). There are also reports of significant interactions with smoking and polymorphisms in carcinogen metabolism genes NAT2 [43] and CYP1A1 [44, 45] and breast cancer susceptibility SNPs [46, 47]. Moreover, BRCA1 and BRCA2 proteins are involved in the repair of DNA damage and it is therefore possible that BRCA1/2 carriers may be more sensitive to effects of carcinogens in cigarette smoke. Thus, despite the limited and inconsistent literature, is it possible there are gene-smoking interactions in relation to breast cancer risk (as there is, for instance, with bladder cancer [48]) and studies may benefit from focusing on more detailed measures and timing of exposure (e.g. peri-menarcheal smoking or pack-years of use) rather than just ever/never smoking.

As in previous studies we excluded from analysis women with prevalent breast or other malignant cancer [11-13, 15-17, 20, 28, 32] or prevalent in-situ breast cancer [13] at recruitment, restricted the analysis to invasive breast cancer [7-18, 20, 26, 28, 30], and adjusted for menopausal status and BMI [8, 10, 11, 13, 16, 18-20, 26, 30, 31], potential confounding variables that may also be influenced by smoking. There was little scope for bias from unascertained mortality or exits, or erroneous reporting of breast cancer, because follow-up for vital and breast cancer status was obtained for 99% of participants and confirmation of reported breast cancers for over 99%. Our smoking information was gained at recruitment and from follow-up questionnaire six years later, and we were able to update smoking status, so that women who gave up smoking were classified as former smokers from that point in time. Only a small number of other cohort studies [13, 16, 20] have been able to update smoking exposure through follow-up. One limitation of our study is that we have no direct information on passive (second hand) smoking and therefore our risk estimates might be underestimated if never-smokers were exposed to passive smoking and if this exposure affects risk of breast cancer [49].

If our results are not due to chance, residual confounding, or unidentified bias, they suggest certain biologic mechanisms deserve further attention, e.g., those involving exposure at peri-menarcheal ages, and gene-environment interactions, either of which may be the direct result of
chemical carcinogenesis or an indirect consequence on hormonal pathways during this susceptible period of breast development.

CONCLUSIONS

We found that smoking was associated with a modest but significantly increased risk of breast cancer, particularly among those who started at adolescent or peri-menarcheal ages, and the relative risk of breast cancer associated with smoking was significantly greater for women with a family history of the disease.

LIST OF ABBREVIATIONS

Body Mass Index, BMI
Confidence Interval, CI
Estrogen receptor, ER
Hazard Ratio, HR
Inter-Quartile Range, IQR
Menopausal Hormone Therapy, MHT
Oral Contraceptive, OC

DECLARATIONS

Ethics approval and consent to participate:
The study was approved by the South Thames Multicentre Research Ethics Committee (ref: MREC 03/01/014) and participants provided informed consent.

Consent for publication:
Not applicable.

Availability of data and material:
The datasets generated during and/or analysed during the current study are not publicly available due to confidentiality reasons but anonymised versions may be available from the corresponding author on reasonable request.
Competing interests:
The authors declare that they have no competing interests.

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Authors' contributions:
AJS and AA designed and obtained funding for the Generations Study. AJS, MEJ and MJS set up and collected data in the Generations Study. MEJ, MJS and LW collected and prepared data for the analysis. MEJ conducted the analyses and drafted the manuscript. All authors contributed to data interpretation and preparation of the final manuscript. All authors read and approved the final manuscript.

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