

1 **TITLE PAGE**

2 **Smoking and risk of breast cancer in the Generations Study cohort**

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37 **ABSTRACT**

38 **Background:** Plausible biological reasons exist why smoking could affect breast cancer risk but
39 epidemiological evidence is inconsistent.

40 **Methods:** We used serial questionnaire information from the Generations Study cohort (United
41 Kingdom) to estimate hazard ratios (HRs) for breast cancer in relation to smoking adjusted for
42 potentially confounding factors including alcohol intake.

43 **Results:** Among 102,927 women recruited 2003–2013, with 7.7 years average follow-up, 1815
44 developed invasive breast cancer. The HR (reference group: never smoker) was 1.14 (95%
45 confidence interval (CI): 1.03–1.25; $P=0.010$) for ever-smoking, 1.24 (95% CI: 1.08–1.43; $P=0.002$) for
46 starting smoking at ages <17 years, and 1.23 (1.07–1.41; $P=0.004$) for starting smoking 1–4 years
47 after menarche. Breast cancer risk was not statistically associated with interval from initiation of
48 smoking to first birth (P -trend=0.97). Women with a family history of breast cancer (ever smoker vs
49 never smoker HR=1.35; 95% CI: 1.12–1.62; $P=0.002$) had significantly larger HR in relation to ever-
50 smoking (interaction: $P=0.039$) than women without (ever smoker vs never smoker HR=1.07; 95% CI:
51 0.96–1.20; $P=0.22$); the interaction was prominent for age started smoking ($P=0.003$) and starting
52 smoking relative to age at menarche ($P=0.0001$).

53 **Conclusions:** Smoking was associated with a modest but significantly increased risk of breast cancer,
54 particularly among women who started smoking at adolescent or peri-menarcheal ages. The relative
55 risk of breast cancer associated with smoking was greater for women with a family history of the
56 disease.

57 **KEYWORDS:**

58 smoking, breast neoplasms, cohort studies

59 **BACKGROUND**

60 The carcinogenic potential of tobacco smoke is unarguable [1, 2] and there are plausible
61 biological reasons why smoking could affect breast cancer risk [2-5]. Reviews of the association
62 between cigarette smoking and breast cancer up to 2004 did not, however, generally find conclusive
63 evidence for a causal relationship in humans [5-7]. More recent epidemiological analyses have
64 reported modest raised risks with current [8-19] or former [8-15, 20] smoking, but questions remain
65 about the extent to which this association is a consequence of confounding by alcohol use, whether
66 risk is increased if smoking starts in adolescence or before first childbirth, and whether risk is
67 modified by family history of breast cancer [1, 2]. We therefore examined risk of invasive breast
68 cancer in relation to smoking in a large cohort study using detailed questionnaire information at
69 recruitment and during follow-up, with adjustment for alcohol consumption and other potentially
70 confounding factors.

71 **METHODS**

72 The Generations Study is a cohort study of over 113,700 women aged 16 or older from the
73 United Kingdom, from whom questionnaire information and informed consent was gained at
74 recruitment since 2003 [21]. Initial recruits to the cohort were from women involved in the breast
75 cancer charity that funded the study, and women who responded to publicity about the study.
76 Women who joined the study were asked to nominate female friends and family members, who
77 were then contacted about joining the study. This referral method continued with subsequent
78 recruits [21]. The first follow-up questionnaire (2½ years after recruitment) was completed by 99%
79 of non-deceased participants, a second (six years after recruitment) by 96%, and a third (9½ years
80 after recruitment) by 94% (of those recruited long enough ago to have entered this round of follow-
81 up). The study was approved by the South East Multi-Centre Research Ethics Committee.

82 Breast and other cancers occurring in the cohort were identified from recruitment and
83 follow-up questionnaires, spontaneous reports to the study centre, and from 'flagging' (see below)
84 for those lost to questionnaire follow-up. Confirmation of diagnosis was obtained from cancer

85 registries in the United Kingdom, ‘flagging’ at the National Health Service Central Registers (virtually
86 complete registers of the populations of England and Wales, and of Scotland, to which study
87 participants can be linked and on which deaths, cancer registrations, and emigrations are ‘flagged’
88 and then periodically reported to authorized medical researchers), pathology reports, and
89 correspondence with patients’ general practitioners.

90 Information on risk factors for breast cancer was obtained from recruitment and follow-up
91 questionnaires. In relation to smoking, women were asked if they had “ever smoked regularly (i.e.
92 most days for at least 6 months)”, if they still smoked regularly, age started and stopped, and
93 number of cigarettes smoked per day at different periods of their lives (during ages: 16–24, 25–49,
94 50+ years). For analysis, we defined the period of ‘current smoking’ to include both current smokers
95 and the year immediately after stopping, to avoid potential ‘reverse-causation’ bias from women
96 who may have stopped smoking during the work-up to a formal breast cancer diagnosis. For alcohol
97 use we asked women if they had been a regular drinker “in the sense of drinking at least one glass of
98 alcohol per week on average”, ages started and stopped, and quantity consumed at different periods
99 of life (during ages: 18–24, 25–49, 50+ years). We converted the quantity of alcohol consumed at
100 each period of life into daily grams of alcohol. We split into three groups the women who reported
101 current drinking (<60g/day, 60+g/day, and amount unknown), and we classified women who had
102 reported stopping drinking as former drinkers. For some women we did not know their current
103 drinking status during follow-up, but we knew they had consumed alcohol in the past and these
104 women were classified as ‘ever-drinkers’. Because we had collected ages or dates at which certain
105 events or changes in lifestyle occurred we were able to update smoking status, alcohol use, parity,
106 oral contraceptive (OC) use, menopausal hormone therapy (MHT) use, and menopausal status, at
107 the ages these episodes occurred through to the second follow-up questionnaire. We updated
108 duration of smoking for current smokers, and time since cessation for former smokers, in yearly
109 increments, using smoking start and stop ages from the recruitment and second follow-up

110 questionnaire. We updated cigarettes smoked per day, pack-years smoked, alcohol consumption,
111 and post-menopausal body mass index (BMI), at the date of the second follow-up questionnaire.

112 *Statistical analysis*

113 The current analytic cohort is based on all women who were recruited to the study during
114 June 2003–December 2013 without prior invasive or *in-situ* breast cancer or other malignant cancer
115 (except non-melanoma skin cancer), or prior mastectomy. The recruitment cut-off at December
116 2013 was selected because at the time of analysis the second follow-up was practically complete for
117 this group of recruits, two-thirds of the cohort had reached the third follow-up, and we had
118 ‘flagging’ information to June 2017. Women entered risk at their date of recruitment and were
119 censored at the earliest date of: invasive breast cancer or *in-situ* breast cancer; other malignancy
120 (except non-melanoma skin cancer); death; most recent follow-up questionnaire (depending on
121 date of recruitment) if completed, or the date most recent follow-up questionnaire was due if cancer
122 and vital status was known from ‘flagging’; or previously completed questionnaire if lost to follow-
123 up. We censored follow-up at *in-situ* breast cancer or other malignancy because we reasoned that
124 if smoking is related to risk of *in-situ* breast cancer or other malignancy, and ensuing treatments or
125 their consequences alter risk of subsequent invasive breast cancer, including subsequent follow-up
126 may obscure associations between smoking and invasive breast cancer.

127 Left-truncated and right censored Cox proportional hazards regression [22] using attained
128 age as the implicit time scale was used to estimate hazard ratios (HR) and 95% confidence intervals
129 (CI) for smoking and risk of first invasive breast cancer. We adjusted for: time since recruitment to
130 cohort (0, 1–2, 3+ years); birth cohort (1908–39, 1940–49, 1950–59, 1960–69, 1970–96); benign
131 breast disease (yes, no); family history of breast cancer in 1st degree relatives (yes, no); socio-
132 economic score (ACORN score as trend, or missing indicator); age at menarche (trend, or missing
133 indicator); age at first pregnancy (trend, or missing indicator); parity (trend, or missing indicator);
134 duration of breastfeeding (trend, or missing indicator); current OC use during follow-up, before

135 menopause (yes, no); alcohol consumption (trend for current drinker 1– <60g/day, indicator
136 variables for never regular, current drinker 60+g/day, past drinker, drinker with unknown details);
137 physical activity (log(metabolic equivalent) trend, missing indicator); pre-menopausal BMI at age 20
138 years (trend, or missing indicator); post-menopausal BMI (trend, or missing indicator); MHT use
139 (never used, ex-user, current estrogen only user, current estrogen plus progestogen user, current
140 user of other types, missing indicator); menopausal status (pre- or post-menopausal) and age at
141 menopause (trend, or missing indicator). BMI was used to create two separate variables: pre-
142 menopausal BMI (potentially available for all women) and post-menopausal BMI (only available at
143 post-menopausal ages). We used BMI at age 20 to represent pre-menopausal BMI. Separately, if a
144 woman was post-menopausal at entry to the cohort we used her BMI at entry for her post-
145 menopausal BMI (and if she was pre-menopausal at this time, post-menopausal BMI was unknown).
146 If a woman was post-menopausal at the time of the follow-up questionnaire we updated from this
147 point in time her post-menopausal BMI with the value from this follow-up questionnaire. Statistical
148 trends were evaluated using continuous values, except for duration and time since cessation of
149 smoking which were based on discrete time-varying annually updated values. For trend analyses
150 where there was an unexposed group (e.g. never smokers in analyses of smoking duration) the
151 unexposed group was not assigned a zero magnitude but was treated as a separate categorical term,
152 as was any missing value group. In particular we adjusted our analyses of smoking and breast cancer
153 for alcohol using daily current alcohol consumption as a continuous measure if within the range 1–
154 <60g/day, and categorical terms for non-drinkers, for those with consumption 60+g/day (because
155 we did not want a minority of women who reported very high consumption to influence unduly the
156 trend with daily consumption), past drinkers, and those for whom details of consumption were
157 missing, by fitting appropriate interaction terms in the Cox regression model. Heterogeneity in HRs
158 by sub-type of breast cancer defined by estrogen receptor (ER) status or morphology was assessed
159 using a data augmentation method [23] and Wald chi-square tests [24]. All statistical tests were
160 two-sided and analyses were conducted using Stata/IC version 14.0 [25].

161 **RESULTS**

162 During 2003–2013 a recruitment questionnaire was completed by 102,940 women who had
163 no previous invasive or *in-situ* breast cancer or other malignancy (except non-melanoma skin
164 cancer). At censoring date 1.1% of women had died. Of the remainder, cancer and vital status was
165 known for 96.5% who had completed the relevant follow-up questionnaire, and a further 2.4% from
166 ‘flagging’ at the National Health Service Central Registers. The remaining 1.1% were lost to follow-
167 up at an earlier date. Thirteen women (including one with breast cancer) were excluded from
168 subsequent analyses because of self-contradictory information for parity or smoking, leaving
169 102,927.

170 Table 1 presents descriptive characteristics at recruitment of the cohort eligible for analysis.
171 The median age at recruitment was 47 years (Inter-Quartile Range (IQR): 36–57). A majority of
172 participants, 64.1%, reported never smoking but only 10.3% were never-regular consumers of
173 alcohol. In relation to alcohol consumption, 12.5% of never-smokers were non-drinkers in contrast
174 to 6.4% of ever-smokers. Among those who reported drinking <60g/day the median alcohol
175 consumption (g/day) was 14.2 (IQR: 8.7–22.1) among never smokers and 19.0 (IQR: 11.9–29.2)
176 among ever smokers. Supplementary Table 1 provides further descriptive characteristics of the
177 cohort in relation to age started smoking, thelarche, parity, menopausal status, and BMI.

178 [TABLE 1 here]

179 During 788,361 person-years (median 6.6 years; mean 7.7 years) of follow-up 1815 invasive
180 breast cancers were diagnosed, of which 1813 were confirmed through national cancer registration
181 or medical records, and the remaining two were self-reported with treatments that imply breast
182 cancer. ER-status data were available for 99.3%, and of these 83.7% were ER-positive. Invasive
183 ductal carcinoma accounted for 78.8%, and lobular 16.4%, of tumours. Further descriptive
184 characteristics of the breast cancer cases are given in Supplementary Table 2.

185 The HR for invasive breast cancer in relation to ever smoking was 1.17 (95% CI: 1.07–1.29;
186 $P=0.0009$) when adjusted only for attained age, 1.13 (95% CI: 1.03–1.24; $P=0.012$) when also
187 adjusted for alcohol consumption, and 1.14 (95% CI: 1.03–1.25; $P=0.010$) when further adjusted for
188 other potentially confounding variables (see Methods and Table 2). All subsequent results are
189 adjusted for attained age, alcohol consumption and the potentially confounding variables, unless
190 otherwise stated.

191 Table 2 presents results for breast cancer overall and by ER status. The HR for ever-smoking
192 was raised for ER-positive (HR=1.12; 95% CI: 1.01–1.24; $P=0.035$) and ER-negative (HR=1.25; 95% CI:
193 0.99–1.58; $P=0.063$) breast cancer, and the difference between the HRs was not significant ($P=0.40$).
194 Breast cancer risk increased significantly with number of cigarettes smoked per day for all breast
195 cancer (P -trend=0.0060) and for ER-positive tumours (P -trend=0.023). Breast cancer risks were
196 raised significantly after 10+ years duration of smoking (10+ years vs never-smoking: $P=0.0004$).
197 Breast cancer risks did not further rise beyond 10 years duration and because of this non-linear
198 relationship there was no significant linear trend with duration of smoking (P -trend=0.24), nor was
199 there significant heterogeneity in the trend by ER status. Pack-years of smoking was associated with
200 breast cancer risk overall (P -trend=0.0069) and ER-positive breast cancer (P -trend=0.024) but not for
201 ER-negative (P -trend=0.16) tumours; there was no significant heterogeneity of the pack-years trend
202 by ER status ($P=0.66$).

203 [TABLE 2 here]

204 The HR within the year after smoking cessation was 2.68 (95% CI: 1.60–4.46), based on 15
205 cases, but for reasons described in Methods this risk period was assigned for further analyses to the
206 ‘current-smoker’ group. On this basis risk of breast cancer was raised in current (HR=1.12; 95% CI:
207 0.89–1.39; $P=0.34$) and former (HR=1.14; 95% CI: 1.03–1.26; $P=0.011$) smokers although only the
208 latter reached statistical significance; there was no significant heterogeneity by ER status. Breast
209 cancer risks were significantly raised within the first 20 years after cessation of smoking and

210 decreased with greater time since cessation although the trend was not significant (P -trend=0.071)
211 and there was no significant heterogeneity in this trend by ER status.

212 There was significant variation in risk of breast cancer by age at start of smoking (Table 3) (P -
213 heterogeneity=0.018; not presented in Table 3). Breast cancer risk was significantly increased if
214 smoking started at ages <17 (HR= 1.24; 95% CI: 1.08–1.43; P =0.0023) or 17–19 (HR= 1.15; 95% CI:
215 1.01–1.31; P =0.030) years relative to non-smokers, but not if it started at older ages. The risk was
216 significantly increased for ER-positive, only for smokers starting at ages <17 years, and no significant
217 risk increase was noted for ER-negative breast cancer. When adjusted for pack-years the breast
218 cancer risk for starting smoking at ages <17 years was (HR= 1.12; 95% CI: 0.96–1.32; P =0.14), and
219 when adjusted for duration of smoking it was (HR=1.16; 95% CI: 0.96–1.40; P =0.11) (not presented in
220 Table 3).

221 [TABLE 3 here]

222 In our questionnaire we asked women only about the amount they smoked per day
223 beginning at age 16; therefore we could not examine smoking intensity at younger ages. There was
224 no significant trend in breast cancer risk, however, in relation to cigarettes smoked per day at ages
225 16–24 years. Relative to age at menarche, breast cancer risks were highest if smoking started at or
226 before menarche (HR=1.40; 95% CI: 0.98–1.99; P =0.061) or 1–4 years after (HR=1.23; 95% CI:
227 1.07–1.41; P =0.0040), with a significant downward trend in breast cancer risk with increasing
228 interval from age at menarche to age at starting smoking (P =0.031). A similar pattern was seen for
229 ER-positive, but was less clear for ER-negative, breast cancer. A weaker relationship was seen with
230 age at thelarche (e.g. 1–4 years after thelarche (HR=1.17; 95% CI: 1.00–1.37; P =0.056)). When
231 adjusted for pack-years of smoking the HRs for age started smoking 1-4 years after menarche
232 (HR=1.12; 95% CI: 0.96–1.31; P =0.15) or thelarche (HR=1.05; 95% CI: 0.88–1.25; P =0.59) were
233 attenuated (not presented in Table 3). There was a comparable attenuation after adjusting for
234 duration of smoking. Among parous women there was a significant trend in breast cancer risk with

235 interval from starting smoking to birth of first child (P -trend=0.013); for an interval of 15+ years the
236 HR was 1.46 (95% CI:1.18–1.81; P =0.0005). However, these results were not adjusted for age at first
237 child birth and parity (not in Tables), and when we adjusted (as shown in Table 3) there were no
238 significantly raised HRs, or trends for all breast cancer or by ER status.

239 When analysed by morphological type (Supplementary Table 3) we found significant
240 associations for ductal breast cancer similar to the results for breast cancer overall, and generally
241 non-significant results for lobular breast cancer. There were no significant interactions by
242 morphological type in the risk of breast cancer with smoking.

243 There was no raised risk of breast cancer with ever-smoking in non-drinkers (HR=0.97; 95%
244 CI: 0.61–1.52; P =0.89) but a significantly raised breast cancer risk in those who had ever been
245 drinkers (HR=1.18; 95% CI: 1.07–1.30; P =0.0010) although the difference in HRs was not significant
246 (P -interaction=0.41) (Table 4). When further stratified by amount of alcohol consumed the HRs for
247 ever-smoking among current drinkers remained raised. Results were similar when we examined
248 breast cancer risk by drinking status for former smokers relative to never smokers (Supplementary
249 Table 4).

250 [TABLE 4 here]

251 We examined further potential risk factor interactions with smoking but found no significant
252 interactions with parity (P =0.095) although for nulliparous ever smoking women there was a
253 statistically significantly increased risk of breast cancer (P =0.012) (Supplementary Table 5), or
254 menopausal status (P =0.73) (Supplementary Table 6), although while the hazard ratio of pre-
255 menopausal ever smokers was somewhat larger than for post-menopausal ever smokers, the former
256 did not reach statistical significance (P =0.088), while the latter did (P =0.040). Nor did we find
257 significant interactions with birth cohort (P =0.092), or BMI at age 20 (P =0.55) or post-menopausal
258 ages (P =0.26), but we did see a significant interaction with family history of breast cancer (P =0.038).
259 There were significant interactions between family history and age at starting smoking (P =0.0029)

260 and starting smoking relative to age at menarche ($P=0.0001$) in relation to risk of breast cancer
261 (Table 5). In particular, among women with a family history of breast cancer, HRs were raised if
262 smoking started at age 20+ years (HR=1.56; 95% CI: 1.17–2.10; $P=0.0028$) or <20 years (HR=1.26;
263 95% CI: 1.02–1.56; $P=0.029$), and if started 5+ years after menarche (HR=1.53; 95% CI: 1.22–1.91;
264 $P=0.0002$), and we note these were somewhat different to the results among women without a
265 family history of breast cancer.

266 [TABLE 5 here]

267 **DISCUSSION**

268 In the Generations Study cohort we found significant but modestly raised risk of invasive
269 breast cancer in ever and former smokers, in women who smoked more than five cigarettes a day,
270 had 10+ pack-years of use, or had stopped for less than 20 years. Some previous studies have
271 reported similar associations with smoking [8-17, 20], cigarettes per day [9-11, 19], pack-years [9-13,
272 17-19, 26-29], and cessation [8, 12, 19, 26, 28], but not all studies find these associations [10, 11, 13,
273 15-17, 19, 20, 29, 30]. We saw significantly raised risk with 10+ years duration of smoking, but no
274 increasing trend beyond 10+ years. Increased risks at long durations (or significant trends) have
275 previously been reported in some studies [8-13, 18-20, 26-28], although some classified non-
276 smokers as smokers with zero duration [12, 20, 26, 28] and this may artefactually produce a
277 significant trend which partly or wholly reflects the difference in risk between non-smokers and
278 smokers (but this may not be the only reason for an association with 20+ years (long duration) of
279 smoking).

280 We found risk was significantly raised in former smokers, as has been previously reported [8-
281 15, 20]. Risk was also raised with current smoking but the numbers of current smokers in our cohort
282 was small and this result did not reach statistical significance, although some other studies have
283 reported significantly raised risks in this group [8-19]. The raised risks for current and former
284 smokers were similar (HR 1.12 and 1.14) and the confidence intervals overlapped, suggesting, within

285 our cohort, no material difference between current and former smokers in relation to breast cancer
286 risk.

287 **Breast cancer sub-types.** We found significant raised risks for ER-positive and ductal breast
288 cancer, which were the most common types in our study, but no significant heterogeneity by ER-
289 status or morphological type of the breast cancer in relation to smoking. The statistical power to
290 examine differences by ER-status or morphology was low in our cohort because of the relative
291 uncommonness of ER-negative and non-ductal type tumours. Some studies have tended to find
292 stronger risks for ER-positive breast cancer [12, 16, 20, 31] but none have found significant
293 interactions and the literature is inconclusive [2]. We observed larger HRs for smoking and pre-
294 menopausal, relative to post-menopausal, breast cancer but the former did not reach statistical
295 significance, and although the literature is variable it does in general suggest a greater relative risk
296 among pre-menopausal women [1, 2]. However, we found no evidence for a significant interaction
297 with menopausal status, similar to other studies [8, 11, 32].

298 **Confounding by alcohol.** Alcohol consumption was associated with smoking and is itself a
299 known risk factor for breast cancer [7]. We adjusted for alcohol intake and although this reduced
300 the strength of the association between smoking and breast cancer (from HR=1.17 to 1.14) the
301 association remained raised and significant. There is, however, concern that statistical adjustment
302 using self-reported alcohol consumption may not be adequate to control fully for confounding by
303 alcohol [7] so to explore further the extent of potential confounding we stratified by alcohol
304 consumption (Table 4). Within each stratum of consumption (<20g/day, 20–40g/day, and 40–
305 <60g/day) the difference in self-reported alcohol intake between never and ever smokers was
306 ≈1g/day, and we calculate this difference in consumption would be associated with <1% change in
307 relative risk of breast cancer (using the alcohol-breast cancer estimate of relative risk from a large
308 collaborative re-analysis [7]). Within each of these strata it would require ever-smokers to be
309 drinking 20g/day more than never-smokers to produce a difference of ≈15% (similar to the 12–17%

310 we saw). This implies the association we observed between ever-smoking and breast cancer may be
311 too large to be explained by differences in alcohol intake alone.

312 We saw no significant association between smoking and breast cancer risk among non-
313 drinkers, in concordance with a collaborative re-analysis of 43 case-control and 10 cohort studies [7],
314 the American Cancer Society's CPS II cohort [16], and a subsequent pooled analysis of 14 cohort
315 studies [8]. It is possible there may be synergistic interaction between ever-smoking and alcohol
316 consumption, and risk of breast cancer, although only one study has reported the interaction as
317 statistically significant [8]. There is some precedent to invoke synergism between smoking and
318 alcohol because, for example, there is an established positive interaction between these two
319 exposures and the aetiology of head and neck cancers [33]. However, non-drinking may occur for
320 cultural or religious reasons, or because of underlying illness or other health issues, and in the UK at
321 least non-drinkers are a minority group; therefore this potential interaction could be a reflection of a
322 particular distribution of breast cancer risk factors among non-drinkers (and inadequate control for
323 confounding among drinkers). Conversely, three other cohort studies found significantly raised risk
324 among non-drinkers [18, 26, 29], although in two the increased raised risks were only in subgroups
325 [26, 29].

326 **Smoking in adolescence.** Based on epidemiological considerations and animal studies the
327 period from puberty to first birth may represent a window of particular susceptibility to breast
328 cancer [34-37]. At puberty the breast is made up of mainly undifferentiated terminal ductal and
329 lobular structures which animal studies show are sensitive to chemical carcinogenesis [34]. At these
330 young ages ionizing radiation exposure also increases risk of breast cancer [37], especially if
331 exposure is within six months of menarche [38]. We found risk of breast cancer in ever-smokers was
332 greatest if smoking started at ages <17 years, or started at peri-menarcheal or, more weakly, peri-
333 thelarcheal ages. A number of other studies have also found raised risks if smoking started in
334 adolescence [8-13, 16-18, 20, 26, 28, 29, 32] or around menarche [11, 16, 26]. However, when we
335 adjusted for pack-years of smoking the raised risks for starting smoking close to age at menarche or

336 thelarche were somewhat attenuated suggesting over-adjustment (because of possible correlation
337 between age starting smoking and pack-years) or confounding by pack-years. Previous studies have
338 not made this adjustment so the relative importance of early initiation or pack-years of use remains
339 unclear.

340 **Smoking before first childbirth** . Young age at first birth and increasing parity confer long-
341 term protection against breast cancer [34, 35] and animal models point to terminal differentiation of
342 breast tissue at full term pregnancy being important in this process [34-36]. Increased risks have
343 been reported for invasive breast cancer if smoking started before first childbirth [8-11, 16, 17, 20,
344 26, 28, 29, 32] but we found the association was only significant if we did not adjust for age at first
345 pregnancy. A number of previous studies have adjusted for age at first pregnancy and still found
346 significant associations with interval to first birth [8, 9, 11-13, 16-18, 20, 26, 28, 29] however it is
347 difficult to determine the adequacy of adjustment. For example, in a large pooled analysis of 14
348 cohort studies there was a strong trend with smoking interval before first birth after adjustment for
349 potential confounders that included age at first birth and number of live births ($P=0.0000002$)
350 whereas after stratification by age at first birth the trends in each strata were weaker ($P=0.12, 0.02,$
351 and 0.28) [8], which is suggestive of confounding.

352 **Interaction with family history**. We found the association between smoking and breast
353 cancer was significantly larger among women with a family history of the disease than those
354 without. Five previous studies have reported on this interaction with family history. Two studies
355 reported no significant interaction but did not present stratified results so we cannot determine if
356 the direction of interaction support or contradict our finding [16, 19]. Three studies reported
357 significant interactions, with one showing increased breast cancer risk with smoking only among
358 those with a positive family history [39], whereas two found breast cancer risk was raised only
359 among those with no family history [15, 18]. Increased risk of breast cancer with smoking has also
360 been seen in some [40, 41], but not all (see review [1] and a large meta-analysis [41]), studies of
361 *BRCA1/2* carriers (or by proxy, women with three or more first degree relatives with breast or

362 ovarian cancer[42]). There are also reports of significant interactions with smoking and
363 polymorphisms in carcinogen metabolism genes *NAT2* [43] and *CYP1A1* [44, 45] and breast cancer
364 susceptibility SNPs [46, 47]. Moreover, BRCA1 and BRCA2 proteins are involved in the repair of DNA
365 damage and it is therefore possible that BRCA1/2 carriers may be more sensitive to effects of
366 carcinogens in cigarette smoke. Thus, despite the limited and inconsistent literature, is it possible
367 there are gene-smoking interactions in relation to breast cancer risk (as there is, for instance, with
368 bladder cancer [48]) and studies may benefit from focusing on more detailed measures and timing of
369 exposure (e.g. peri-menarcheal smoking or pack-years of use) rather than just ever/never smoking.

370 As in previous studies we excluded from analysis women with prevalent breast or other
371 malignant cancer [11-13, 15-17, 20, 28, 32] or prevalent *in-situ* breast cancer [13] at recruitment,
372 restricted the analysis to invasive breast cancer [7-18, 20, 26, 28, 30], and adjusted for menopausal
373 status and BMI [8, 10, 11, 13, 16, 18-20, 26, 30, 31], potential confounding variables that may also be
374 influenced by smoking. There was little scope for bias from unascertained mortality or exits, or
375 erroneous reporting of breast cancer, because follow-up for vital and breast cancer status was
376 obtained for 99% of participants and confirmation of reported breast cancers for over 99%. Our
377 smoking information was gained at recruitment and from follow-up questionnaire six years later,
378 and we were able to update smoking status, so that women who gave up smoking were classified as
379 former smokers from that point in time. Only a small number of other cohort studies [13, 16, 20]
380 have been able to update smoking exposure through follow-up. One limitation of our study is that
381 we have no direct information on passive (second hand) smoking and therefore our risk estimates
382 might be underestimated if never-smokers were exposed to passive smoking and if this exposure
383 affects risk of breast cancer [49].

384 If our results are not due to chance, residual confounding, or unidentified bias, they suggest
385 certain biologic mechanisms deserve further attention, e.g., those involving exposure at peri-
386 menarcheal ages, and gene-environment interactions, either of which may be the direct result of

387 chemical carcinogenesis or an indirect consequence on hormonal pathways during this susceptible
388 period of breast development.

389 **CONCLUSIONS**

390 We found that smoking was associated with a modest but significantly increased risk of
391 breast cancer, particularly among those who started at adolescent or peri-menarcheal ages, and the
392 relative risk of breast cancer associated with smoking was significantly greater for women with a
393 family history of the disease.

394 **LIST OF ABBREVIATIONS**

395 Body Mass Index, BMI

396 Confidence Interval, CI

397 Estrogen receptor, ER

398 Hazard Ratio, HR

399 Inter-Quartile Range, IQR

400 Menopausal Hormone Therapy, MHT

401 Oral Contraceptive, OC

402 **DECLARATIONS**

403 **Ethics approval and consent to participate:**

404 The study was approved by the South Thames Multicentre Research Ethics Committee (ref: MREC
405 03/01/014) and participants provided informed consent.

406 **Consent for publication:**

407 Not applicable.

408 **Availability of data and material:**

409 The datasets generated during and/or analysed during the current study are not publicly available
410 due to confidentiality reasons but anonymised versions may be available from the corresponding
411 author on reasonable request.

412 **Competing interests:**

413 The authors declare that they have no competing interests.

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419 writing or approving the manuscript.

420 **Authors' contributions:**

421 AJS and AA designed and obtained funding for the Generations Study. AJS, MEJ and MJS set up and
422 collected data in the Generations Study. MEJ, MJS and LW collected and prepared data for the
423 analysis. MEJ conducted the analyses and drafted the manuscript. All authors contributed to data
424 interpretation and preparation of the final manuscript. All authors read and approved the final
425 manuscript.

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